Ethnic variation in the conversion of prediabetes to diabetes among immigrant populations relative to Canadian-born residents: a population-based cohort study

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ABSTRACT

Objective The aim of this study was to compare absolute and relative rates of conversion from prediabetes to diabetes among non-European immigrants to Europeans and Canadian-born residents, overall, and by age and level of glycemia.

Research design and methods We conducted a retrospective cohort population-based study using administrative health databases from Ontario, Canada, to identify immigrants (n=23465) and Canadian born (n=111085) aged ≥20 years with prediabetes based on laboratory tests conducted between 2002 and 2011. Individuals were followed until 31 December 2013 for the development of diabetes using a validated algorithm. Immigration data was used to assign ethnicity based on country of origin, mother tongue, and surname. Fine and Gray’s survival models were used to compare diabetes incidence across ethnic groups overall and by age and glucose category.

Results Over a median follow-up of 5.2 years, 8186 immigrants and 39722 Canadian-born residents developed diabetes (7.1 vs 6.1 per 100 person-years, respectively). High-risk immigrant populations such as South Asians (HR: 1.72, 95% CI 1.55 to 1.99) and Southeast Asians (HR: 1.65, 95% CI 1.46 to 1.86) had highest risk of converting to diabetes compared with Western Europeans (referent). Among immigrants aged 20–34 years, the adjusted cumulative incidence ranged from 18.4% among Eastern Europeans to 52.3% among Southeast Asians. Conversion rates increased with age in all groups but were consistently high among South Asians, Southeast Asians and Sub-Saharan African/Caribbeans after the age of 35 years. On average, South Asians converted to diabetes 3.1–4.6 years earlier than Western Europeans and at an equivalent rate of conversion to Western Europeans who had a 0.5 mmol/L higher baseline fasting glucose value.

Conclusions High-risk ethnic groups converted to diabetes more rapidly, at younger ages, and at lower fasting glucose values than European populations, leading to a shorter window for diabetes prevention.

INTRODUCTION

Diabetes is one of the fastest growing epidemics worldwide. Widespread shifts in lifestyle and aging of the world’s populations have resulted in an unprecedented rise in the global prevalence of diabetes and its complications.1,2 In recent decades, rapid shifts in urbanization caused by economic growth and rural to urban migration have fueled rising rates of diabetes in low-income and middle-income countries by increasing access to unhealthy food options and reducing the opportunities local residents have for physical activity.1–4 According to the International Diabetes Federation, 693 million people are expected to have diabetes by 2045, with the
largest increases expected in India and China. Based on current estimates, diabetes prevalence is highest in regions within the Western Pacific, Middle East, North Africa and Central America, with some countries reporting a prevalence of 15%–20% or higher among adults aged 20–79 years. 

Ethnicity is an important factor influencing the risk of metabolic disease. Approximately 5%–15% of people with prediabetes convert to diabetes each year depending on the underlying characteristics of the population (eg, ethnicity) and the type and severity of their glucose abnormality. The landmark US Diabetes Prevention Program (DPP) reported diabetes incidence among participants who were enrolled in the placebo arm, finding 25% higher rates among African-American, Hispanic, American Indian and Asian participants than non-Hispanic whites. Several other studies have demonstrated high rates of conversion within specific ethnic populations including Chinese, Japanese, European and South Asians, among others, even after accounting for individual and socioeconomic factors. However, comparisons across a broader range of ethnicities is lacking.

There is mounting evidence that intensive lifestyle interventions can substantially reduce the rate of conversion from prediabetes to diabetes. Clinical practice guidelines recommend earlier and more frequent screening for diabetes in high-risk populations, including those of non-European heritage, to avoid missed opportunities for diabetes prevention. However, it is unclear whether the window to incorporate diabetes prevention strategies is similar across ethnic groups in whom a diagnosis of prediabetes has been made.

The primary aim of this study was to examine the absolute and relative risk of conversion from prediabetes to type 2 diabetes in non-European ethnic groups relative to Europeans and Canadian-born residents of the same age. The secondary aim of this study was to derive a similar follow-up time and level of glycemia among non-Europeans that conferred an equivalent probability of converting to type 2 diabetes as Western Europeans. We hypothesized that non-European populations, particularly South Asians, would have a substantially higher risk of converting to diabetes than that of Europeans and that compared with the latter, non-Europeans would convert to diabetes more rapidly, at younger ages, and at lower levels of glycemia.

METHODS

Study design and population

We used population-level health and laboratory data from Ontario, Canada, to identify adults who were found to have prediabetes on laboratory testing between 2002 and 2011. Individuals were then followed until 31 December 2013 for the development of diabetes. Our primary population of interest were immigrants to Canada because information on ethnicity was available through federal immigration data for those having immigrated since 1985. We also ascertained data on long-term residents with prediabetes who did not meet this criteria. This comparison population largely included individuals who were Canadian born and those who immigrated to Canada prior to 1985 (referred to herein as ‘Canadian born’). Nearly 30% of Ontario’s current residents were born outside of Canada, enabling us to assemble a large multiethnic cohort of adults with prediabetes and to follow their health outcomes over time. Canada’s healthcare system is designed to provide universal and comprehensive coverage of healthcare services to all permanent residents. As such, Ontario’s administrative healthcare data capture service utilization and outcomes for all permanent residents regardless of whether they are Canadian born or landed immigrants. For this study, we combined a variety of health and non-healthcare datasets, including detailed records at the time of immigration. Records were linked across databases using unique, encoded identifiers and analyzed at ICES, a research institute affiliated with the University of Toronto.

The study included data on all adults, aged 20–85 years, who met the case definition of prediabetes based on glucose tests performed between 1 January 2002 and 31 December 2011 at one of the province’s 225 Dynacare Medical Laboratory (DML) locations. DML is one of the largest commercial laboratories in Ontario, performing 25%–30% of all glucose tests ordered by Ontario doctors. These are directly reimbursed through the province’s healthcare system, and prior research suggests that the demographic characteristics of individuals who are tested through DML are similar to those tested elsewhere.

Prediabetes was defined as impaired fasting glucose (IFG; 6.1–6.9 mmol/L), impaired glucose tolerance (IGT; 2-hour glucose of 7.8–11.0 mmol/L after a 75 g oral glucose tolerance test (OGTT) or glycosylated hemoglobin (HbA1c) of 6.0%–6.4% using criteria established by the WHO and Diabetes Canada. A 2-year look-back window was set to ensure that individuals did not have blood glucose values in the diabetes range prior to meeting the case definition for prediabetes.

Measures

Ethnicity and immigration status were the main exposures in this study. The Immigration, Refugees and Citizenship Canada Permanent Resident (IRCC-PR) database was used to classify people who landed in Canada after 1985 as immigrants. All others were classified as Canadian-born residents. Among immigrants, a validated algorithm was used to derive ethnic groups based on country of birth, mother tongue, and surname. As a result, people of different ethnic origins were classified into the following groups: Eastern European, Western European, Latin American, sub-Saharan African and Caribbean, West Asian and Arab, East Asian, Southeast Asian, and South Asian (online supplementary table 1). Immigration-related information such as country of birth, mother tongue, education level, immigration visa category, marital status, landing date and English language proficiency was obtained from the IRCC-PR dataset. Demographic and vital statistics information was obtained from Ontario’s Registered Persons...
Database (RPDB). The IRCC-PR and the RPDB have been probabilistically linked with a linkage rate of 86.4%. Since information on individual income was not available for our cohort, neighborhood-level income quintiles were assigned instead using information collected in the 2006 Canadian Census. These values were based on small geographical units defined by Statistics Canada, dissemination areas, and adjusted for household and community size. Dissemination areas consist of one or more city blocks and a population size of 400–700 persons.

Individuals were followed for a minimum of 2 years until 31 December 2013 to identify new diabetes cases using the Ontario Diabetes Database (ODD). The ODD uses a validated algorithm based on hospitalization and physicians’ services claims data to identify individuals with diabetes (excluding those with gestational diabetes) and has a sensitivity of 89% and a specificity of over 98%. Individuals with a record in the ODD or with any diabetes-range glucose measures in DML prior to baseline were excluded from the study cohort.

**Statistical analyses**

Crude diabetes incidence rates were calculated per 100 person-years with the number of people who developed diabetes on follow-up as the numerator and total person-years as the denominator. Unadjusted and adjusted cumulative incidence function curves for diabetes incidence were generated for immigrants of different ethnic origins using Fine and Gray’s survival modeling. Using Fine and Gray’s cumulative incidence function with competing risks allowed us to estimate the incidence of diabetes in the presence of a competing event. In this case, death was treated as a competing event in all analyses. In addition, individuals were censored if they lost healthcare coverage (indicating a move out of province). Also, the proportionality of hazards assumption was assessed using log-negative-log of the Kaplan-Meier survival curves against the log of survival time. Also, the assumption of linearity was assessed for continuous variables such as age. All assumptions were met for all covariates in the models.

Among immigrants, we compared the incidence of diabetes in each ethnic group relative to Western Europeans (referent population), adjusting for age and sex. Next, we generated fully adjusted models accounting for additional covariates including area income, education, immigration visa category, marital status, and duration in Canada. We conducted stratified analyses by age group and sex to compare the effects of ethnicity on diabetes incidence across these groups. Similar analyses were conducted to compare the risk of conversion to diabetes among immigrants (all ethnicities combined) and Canadian-born residents with prediabetes at baseline.

As a secondary analysis, we compared the rate of conversion from prediabetes to diabetes among immigrants in each ethnic group using three different thresholds of impaired fasting glucose to define prediabetes. The categories included fasting values of: (1) 5.7–6.0 mmol/L, (2) 6.1–6.5 mmol/L, and (3) fasting 6.6–6.9 mmol/L. Virtually all individuals in our cohort had fasting glucose measurements on or near the index date and few had baseline HbA1c or an OGTT performed; therefore, only fasting glucose values were considered in assigning these categories. As above, the cumulative incidence rates were also adjusted for sociodemographic and immigration factors.

In addition, we used cumulative incidence function curves to plot the relation between follow-up time (x) and cumulative incidence probability for the development of diabetes (y). We then used a line of best fit based on a polynomial (quadratic) equation \((y = a + bx + cx^2)\) to derive the follow-up time in years (x) from baseline that corresponded to a cumulative incidence probability (y) of 0.25 or 0.50 (ie, the time required for 25%–50% of the population to develop diabetes). We used a cumulative incidence probability threshold of 0.25 to compare ethnic groups that had baseline fasting glucose concentrations between 6.1 and 6.5 mmol/L, as the overall incidence of diabetes was lower in this group, and a cumulative incidence probability threshold of 0.50 to compare ethnic groups that had baseline fasting glucose concentrations between 6.6 mmol/L and 6.9 mmol/L, as this group had a higher incidence of diabetes. Differences in follow-up time in years were calculated by subtracting the value obtained (above) for each ethnic group from that derived for Western Europeans (the referent population).

Lastly, we ran additional models to examine whether there was an interaction between ethnicity and duration in Canada (for immigrants only) to test for any modifying effects of the amount of time spent in Canada on the conversion to diabetes.

All analyses were conducted using an alpha level of 0.05 and by using SAS Enterprise, V.7.4.

**RESULTS**

Overall, there were 23,465 immigrants and 111,085 Canadian-born residents with prediabetes in our cohort. Baseline characteristics of the study population by immigration status and ethnicity are shown in table 1 and online supplementary table 2. Briefly, immigrants were younger at baseline compared with Canadian-born residents (mean age 49.9 vs 57.1, respectively) and more likely to live in lower income neighborhoods (table 1). More than half of immigrants arrived to Canada with high school education attainment or less, with the exception of Eastern Europeans. Although the majority of immigrant populations landed with an economic visa that includes skilled workers and their family members, a larger proportion of immigrants of sub-Saharan African/Caribbean, Latin American, and West Asians/Arab descent arrived to Canada with a refugee status (online supplementary table 2).

The median follow-up time was 5.2 years, and over this period 8186 immigrants and 39,722 Canadian-born residents progressed from prediabetes to diabetes, respectively. Compared with the latter, the crude incidence of
The non-natives incidence of diabetes was substantially higher among immigrants (7.1 vs 6.1 per 100 person-years) (table 2). These ethnic differences were also apparent among younger adults. Over 12 years, nearly 50% of South Asians and Southeast Asians aged 20–34 converted to diabetes (figure 1). The incidence of diabetes rose further with increasing age reaching a cumulative conversion rate of nearly 60% for adults of South Asian, Southeast Asian, West Asian and Arab, and Sub-Saharan African and Caribbean origins at the age of 35 years and older. For other ethnic groups, such as Latin Americans, the risk of converting to diabetes approached 60% after age 65 years.

Over the study period, the adjusted cumulative incidence of diabetes was increased with higher levels of fasting plasma glucose at baseline. However, ethnic differences in diabetes incidence persisted regardless of baseline glycemia (online supplementary figure 1). Nonetheless, non-Europeans converted to diabetes 3.1–4.6 years earlier than European populations (figure 2).

Adjusted conversion rates among non-Europeans were similar to that of Europeans who had higher levels of glycemia at baseline. For example, the cumulative incidence of diabetes among South Asians who had baseline fasting glucose values of 6.1–6.5 mmol/L was equivalent to that of Western Europeans who had baseline fasting glucose values of 6.6–6.9 mmol/L. In both instances, ~60% of individuals developed diabetes over the study period (figure 3). At lower baseline fasting glucose values of 5.7–6.0 mmol/L, diabetes was higher among immigrants (7.1 vs 6.1 per 100 person-years) (table 2).

### Table 1: Characteristics of study population at baseline by immigration status (n=134 550)

<table>
<thead>
<tr>
<th>Variables</th>
<th>Canadian-born residents n=111 085</th>
<th>Immigrants n=23 465</th>
</tr>
</thead>
<tbody>
<tr>
<td>Mean age±SD*</td>
<td>57.1±12.9</td>
<td>49.9±12.8</td>
</tr>
<tr>
<td>Females</td>
<td>42.7</td>
<td>43.8</td>
</tr>
<tr>
<td>Income quintile †</td>
<td></td>
<td></td>
</tr>
<tr>
<td>Q1 (lowest)</td>
<td>16.7</td>
<td>28.7</td>
</tr>
<tr>
<td>Q2</td>
<td>20.1</td>
<td>23.9</td>
</tr>
<tr>
<td>Q3</td>
<td>20.7</td>
<td>20.2</td>
</tr>
<tr>
<td>Q4</td>
<td>21.0</td>
<td>17.3</td>
</tr>
<tr>
<td>Q5 (highest)</td>
<td>21.4</td>
<td>10.0</td>
</tr>
<tr>
<td>Prediabetes criteria met at baseline for the entire population</td>
<td></td>
<td></td>
</tr>
<tr>
<td>IFG of 6.1–6.9 mmol/L</td>
<td>79.6</td>
<td>62.2</td>
</tr>
<tr>
<td>IGT of 7.8–11.0 mmol/L</td>
<td>0.91</td>
<td>1.30</td>
</tr>
<tr>
<td>HbA1c of 6.0%–6.4%</td>
<td>19.5</td>
<td>36.7</td>
</tr>
</tbody>
</table>

All values expressed as % unless otherwise specified.

*New cases of diabetes were determined based on entry into the Ontario Diabetes Database. Cases censored at time when diabetes was ascertained, lost healthcare coverage, died or 31 December 2013.

†Adjusted for age, sex, area income, education, ethnicity, marital status, immigration visa category, and duration. These covariates were not available for the Canadian-born residents.

§Ethnic groups were derived based on country of birth, mother tongue, and a validated algorithm that identifies ethnic groups based on surnames for South Asian and Chinese populations only.

### Table 2: Association between ethnicity/immigration status and diabetes incidence (n=134 550)

<table>
<thead>
<tr>
<th>Variable</th>
<th>No. new cases of diabetes*</th>
<th>Person-years of follow-up</th>
<th>Crude incidence (no./100 person-years)</th>
<th>Age, sex- adjusted HR (95% CI)*†</th>
<th>Fully adjusted HR (95% CI)*‡</th>
</tr>
</thead>
<tbody>
<tr>
<td>Immigrants, ethnicity§</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>South Asian</td>
<td>2417</td>
<td>28 597</td>
<td>8.5</td>
<td>1.71 (1.55 to 1.89)</td>
<td>1.72 (1.55 to 1.99)</td>
</tr>
<tr>
<td>Southeast Asian</td>
<td>916</td>
<td>11 492</td>
<td>8.0</td>
<td>1.58 (1.41 to 1.77)</td>
<td>1.65 (1.46 to 1.86)</td>
</tr>
<tr>
<td>Sub-Saharan African and Caribbean</td>
<td>890</td>
<td>11 780</td>
<td>7.6</td>
<td>1.55 (1.39 to 1.74)</td>
<td>1.63 (1.45 to 1.83)</td>
</tr>
<tr>
<td>Latin American</td>
<td>702</td>
<td>9 506</td>
<td>7.4</td>
<td>1.49 (1.32 to 1.68)</td>
<td>1.58 (1.41 to 1.78)</td>
</tr>
<tr>
<td>East Asian</td>
<td>1270</td>
<td>18 119</td>
<td>7.0</td>
<td>1.34 (1.23 to 1.52)</td>
<td>1.53 (1.35 to 1.73)</td>
</tr>
<tr>
<td>West Asian and Arab</td>
<td>756</td>
<td>9 271</td>
<td>8.2</td>
<td>1.63 (1.45 to 1.84)</td>
<td>1.37 (1.23 to 1.53)</td>
</tr>
<tr>
<td>Eastern European</td>
<td>769</td>
<td>16 710</td>
<td>4.6</td>
<td>0.92 (0.82 to 1.03)</td>
<td>0.92 (0.82 to 1.04)</td>
</tr>
<tr>
<td>Western European</td>
<td>466</td>
<td>9 510</td>
<td>4.9</td>
<td>Referent</td>
<td>Referent</td>
</tr>
<tr>
<td>Entire cohort, immigration status</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Immigrants to Canada</td>
<td>8186</td>
<td>64 982</td>
<td>7.1</td>
<td>1.18 (1.15 to 1.20)</td>
<td>1.15 (1.12 to 1.18)</td>
</tr>
<tr>
<td>Canadian-born residents</td>
<td>39 722</td>
<td>115 043</td>
<td>6.1</td>
<td>Referent</td>
<td>-</td>
</tr>
</tbody>
</table>

*New cases of diabetes were determined based on entry into the Ontario Diabetes Database. Cases censored at time when diabetes was ascertained, lost healthcare coverage, died or 31 December 2013.

†Adjusted for age, sex, and area income.

‡Among immigrants only, adjusted for age, sex, area income, education, ethnicity, marital status, immigration visa category, and duration. These covariates were not available for the Canadian-born residents.

§Ethnic groups were derived based on country of birth, mother tongue, and a validated algorithm that identifies ethnic groups based on surnames for South Asian and Chinese populations only.
the rate of conversion was 27% among South Asians and 12% among Western Europeans (online supplementary figure 1). Similar findings were apparent for other high-risk non-European groups (data not shown).

Overall, ethnic disparities were consistent across sexes, and in some high-risk ethnic groups, these effects were stronger for men (online supplementary figure 2). Also, duration in Canada modified the relationship between ethnicity and conversion to diabetes such that ethnic disparities that were evident in the first 10 years of living in Canada were more accentuated with greater time (>10 years) since landing (online supplementary figure 3).

**DISCUSSION**

In this study, rates of conversion from prediabetes to diabetes were substantially higher among non-European immigrants to Canada than Europeans regardless of their age, sex, and level of glycemia. Relative to Western Europeans, individuals of South Asian, Southeast Asian, and sub-Saharan African/Caribbean descent had one and a half to twofold higher adjusted conversion rates, with ~60% developing diabetes over 12 years. Moreover, these ethnic differences emerged at young ages. One in two South Asians and Southeast Asians with prediabetes who were between the ages of 20 years and 34 years converted to diabetes over this same period. These higher risk populations also bear an increased incidence of prediabetes, adding to their growing burden of type 2 diabetes in middle age. Our findings suggest that a single definition of prediabetes confers a different prognosis depending on the ethnicity of the population. Moreover, non-Europeans with prediabetes developed diabetes years earlier than Europeans with the same level of glycemia, suggesting that the window to adopt intensive lifestyle strategies is much shorter for high-risk ethnic groups.

This research adds to previous literature examining the rate of conversion from prediabetes to diabetes in various populations. However, our study was the first to provide a detailed analysis of this phenomenon in a large multiethnic sample, allowing for a comparison of conversion rates across immigrants from various world regions. Similar to other studies that have demonstrated a higher prevalence of diabetes among individuals of South Asian heritage, we found much higher rates of conversion to diabetes in South Asian immigrants with prediabetes compared with those of European descent. In this study, the crude incidence of diabetes among South Asians and Southeast Asians (8.6 and 8.0 per 100 person-years, respectively) was comparable with rates reported for Pima Indians and Asian Indians but higher than those reported...
Figure 2  Adjusted cumulative incidence of diabetes among immigrants, by fasting glucose category at baseline. Adjusted for age, sex, ethnicity, area income, education, marital status, immigration visa category, and duration. New cases of diabetes were determined based on entry into the Ontario Diabetes Database. Cases censored at time when diabetes was ascertained, lost health care coverage, died, or 31 December 2013. All lines based on line of best fit using a polynomial equation, $R^2>0.99$ for each.

The risk of diabetes among immigrants may be shaped, in part, by socioeconomic, environmental, and political influences that lead to immigration and acculturation to post immigration life. A significant proportion of immigrants in our study lived in lower income neighborhoods, a factor that was independently associated with a higher risk of diabetes. Furthermore, high-risk immigrant groups may be more likely to live in suburban neighborhoods that offer fewer opportunities for physical activity. These findings and others have led to the theory known as the healthy migrant phenomena. In line with this concept, ethnic differences in the rate of conversion from prediabetes to diabetes in our study were greater among populations who had been in Canada more than 10 years. Differences in acculturation to a western lifestyle may contribute to these findings as some studies suggest that immigrant populations that preserve their cultural traditions with respect to diet and physical activity gain less weight and have more favorable metabolic outcomes.

To our knowledge, this was the first population-based longitudinal cohort analysis comparing the risk of conversion from prediabetes to diabetes across a broad range of ethnic groups living in similar environments. Additionally, access to linked administrative health and immigration data allowed us to study this phenomenon in a large sample and to examine subsets of the population on the basis of age group, age and duration in Canada. However, our study has several limitations. First, although we used validated algorithms to derive ethnic origins, some degree for Native Americans, in general, and Europeans living in Denmark and Spain. The risk of diabetes among immigrants may be shaped, in part, by socioeconomic, environmental, and political influences that lead to immigration and acculturation to post migration life. A significant proportion of immigrants in our study lived in lower income neighborhoods, a factor that was independently associated with a higher risk of diabetes. Furthermore, high-risk immigrant groups may be more likely to live in suburban neighborhoods that offer fewer opportunities for physical activity. These findings and others have led to the theory known as the healthy migrant phenomena. In line with this concept, ethnic differences in the rate of conversion from prediabetes to diabetes in our study were greater among populations who had been in Canada more than 10 years. Differences in acculturation to a western lifestyle may contribute to these findings as some studies suggest that immigrant populations that preserve their cultural traditions with respect to diet and physical activity gain less weight and have more favorable metabolic outcomes.

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Figure 3 Adjusted cumulative incidence function for diabetes among South Asians and Western Europeans, by fasting glucose category. Adjusted for age, sex, ethnicity, area income, education, marital status, immigration visa category and duration. New cases of diabetes were determined based on entry into the Ontario Diabetes Database. Cases censored at time when diabetes was ascertained, lost health care coverage, died, or 31 December 2013. All lines based on line of best fit using a polynomial equation, \( R^2 > 0.99 \) for each.

This study has several policy and practice implications for diabetes prevention. The rapid rate of conversion to diabetes among young adults from high-risk ethnic groups is concerning given the lack of specific guidance on how early to screen these populations for prediabetes. There are missed opportunities for diabetes prevention if high-risk groups with prediabetes are left unidentified and a shorter window to prevent diabetes development once a diagnosis of prediabetes is made. Earlier screening strategies and targeted interventions to promote healthier lifestyles in high-risk communities may help to address these challenges. In particular, community-based programs, including those similar to the DPP-translated interventions, that are culturally sensitive may prove to be more effective in reducing the burden of diabetes. Further research is needed to understand the optimal approaches to screening and prevention in order to reduce the burden of diabetes in high-risk populations.

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of misclassification may have occurred due to the broad categorization of ethnicity based on world regions, since ethnic groups living within a given region may be quite heterogeneous. Second, these data sources lacked information on diet and physical activity patterns, health beliefs, body mass index, family history of diabetes and other markers of insulin resistance, as well as premigration risk factors. Thus, causes of the ethnic differences observed in this study could not be fully elucidated. Third, our definition of Canadian-born residents included immigrants who arrived in Canada prior to 1985, which may lead to some degree of misclassification. However, this would tend to reduce differences between immigrants and Canadian-born residents and thus go against our null hypothesis. Fourth, we used a single commercial laboratory for identifying prediabetes cases, which captures approximately one-third of blood tests in Ontario and therefore did not include all true prediabetes cases in our setting, and repeat laboratory testing at a different laboratory would not have been captured. However, we used a validated algorithm to identify new cases of diabetes across the province, which allowed for comprehensive follow-up data for all individuals in our cohort. Fifth, although we did not measure healthcare utilization (ie, primary care visits), there is no evidence that screening rates are lower among immigrants. Therefore, under Ontario’s universal healthcare plan, primary care use and diabetes screening rates are high among Ontarians.
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Competing interests None declared.

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Ethics approval This study received ethics approval from the University of Toronto and Institutional Review Board at Sunnybrook Health Sciences Centre.

Provenance and peer review Not commissioned; externally peer reviewed.

Data availability statement The data presented in this study are housed in ICES and are not publicly available due to privacy agreements.

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